Prelim Exam

**Linear Models** 

Spring 2014

## Preliminary Examination: LINEAR MODELS

Answer all questions and show all work.

1. Suppose that data  $\{(x_{ij}, y_{ij}) : i = 1, ..., k, j = 1, ..., J\}$  can be modeled as having a common slope and possibly different intercepts using the linear model,

$$Y_{ij} = \beta_i + \gamma x_{ij} + \epsilon_{ij},$$

where  $\{\epsilon_{ij}\}\$  are independently and identically distributed  $N(0, \sigma^2)$  random variables. Assume that no vector  $(x_{i1}, \ldots, x_{iJ})$ , for  $i = 1, \ldots, k$ , is proportional to the vector of 1s.

- a. Determine the ordinary-least-squares estimator of  $(\beta_1, \ldots, \beta_k, \gamma)'$ .
- b. Give an explicit expression for the size  $\alpha$  likelihood-ratio test of the hypothesis,

$$H_0: \beta_1 = \cdots = \beta_k = 0$$
 versus  $H_a:$  not  $H_0$ 

- c. Compute the power of the test that you derived in part (b). (There are several ways of defining the non-centrality parameter for the test. Pick any one of these, and use it consistently in this part.) Show that the power is independent of  $\gamma$ .
- 2. Let  $X_1, \ldots, X_n$  be iid random variables with a  $N(\theta, \sigma^2)$  distribution conditional on  $\theta$  and  $\sigma^2$ , and  $\theta$  and  $\sigma^2$  are unknown.
  - a. Give the classical test statistic for testing  $\sigma^2 = \sigma_0^2$  against  $\sigma^2 \neq \sigma_0^2$ .

Suppose that the following assumptions are made concerning  $\theta$  and  $\sigma^2$ :  $\theta$  follows a  $N(\theta_0, \tau^2)$  distribution with both  $\theta_0^2$  and  $\tau^2$  known, and  $\sigma^2 = \sigma_0^2 > 0$ .

- b. The *(joint) marginal* distribution of  $(X_1, \ldots, X_n)$  (i.e., unconditional on the parameters) is multivariate normal. Derive the mean and variance-covariance matrix for this distribution.
- c. Find the distribution of the random variable

$$T(X_1, \dots, X_n) = \frac{1}{\sigma_0^2} \sum_{i=1}^n (X_i - \bar{X})^2 + \frac{n(\bar{X} - \theta_0)^2}{n\tau^2 + \sigma_0^2}$$

under the marginal distribution of  $(X_1, \ldots, X_n)$ .

d. Let  $(x_1, \ldots, x_n)$  denote the observed values of  $(X_1, \ldots, X_n)$ . The Bayesian prior predictive p-value to validate the specified model is defined as

$$p = P(T(X_1, \ldots, X_n) \ge T(x_1, \ldots, x_n))$$

under the marginal distribution of  $(X_1, \ldots, X_n)$ . Compare this Bayesian prior predictive p-value to the classical p-value for testing  $\sigma^2 = \sigma_0^2$  against  $\sigma^2 \neq \sigma_0^2$  in part (a).

3. Consider a simple linear regression model,

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i, i = 1, \dots, n,$$

where  $\{\epsilon_i\}$  are iid  $N(0, \sigma^2)$  random variables.

Let  $\mathbf{Y}' = (Y_1, Y_2, \dots, Y_n), \, \boldsymbol{\beta}' = (\beta_0, \beta_1), \text{ and } \mathbf{X}' = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ x_1 & x_2 & \cdots & x_n \end{bmatrix}$ . Furthermore, for known constants  $\{z_i : i = 1, \dots, n\}$ , define  $\mathbf{Z}' = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ z_1 & z_2 & \cdots & z_n \end{bmatrix}$ , and assume that  $\mathbf{Z}'\mathbf{X}$  is a non-singular matrix.

- a. Show that the so-called instrumental-variable estimator,  $\tilde{\mathbf{b}} = (\mathbf{Z}'\mathbf{X})^{-1}\mathbf{Z}'\mathbf{Y}$ , is an unbiased estimator of the vector  $\boldsymbol{\beta}$ .
- b. Find the sampling distribution of  $\tilde{\mathbf{b}} = (\tilde{b}_0, \tilde{b}_1)'$ , including its mean vector and variancecovariance matrix. Furthermore, let  $\hat{\mathbf{b}} = (\hat{b}_0, \hat{b}_1)'$  be the OLS estimator of  $\beta$ . Prove that  $var(\tilde{b}_1) \ge var(\hat{b}_1)$ .
- c. Let  $P_Z$  denote the projection matrix on the column space C(Z).
- c-i. Show that  $\mathbf{Y}'(\mathbf{I} \mathbf{P}_{\mathbf{Z}})\mathbf{Y}/\sigma^2$  has a chi-squared distribution. Give its degrees of freedom and the non-centrality parameter.
- c-ii. What is the value of the non-centrality parameter when  $\beta_1 = 0$ ? Justify your answer.
  - d. Let the residuals corresponding to the instrumental-variable estimator  $\tilde{\mathbf{b}}$  be denoted by

$$\tilde{\mathbf{e}} = \mathbf{Y} - \mathbf{X}\mathbf{b}.$$

- d-i. Show that  $\tilde{\mathbf{b}}$  and  $\tilde{\mathbf{e}}$  are independently distributed and give the sampling distribution of  $\tilde{\mathbf{e}}$ .
- d-ii. Now let **Q** denote the  $n \times n$  matrix so that  $\tilde{\mathbf{e}} = \mathbf{Q}\mathbf{Y}$ . Show that the matrix **Q** is not symmetric. Is the matrix **Q** idempotent? Explain why or why not.
- d-iii. Give a necessary and sufficient condition for the distribution of  $\tilde{\mathbf{e}}'\tilde{\mathbf{e}}/\sigma^2$  to be a chi-squared distribution.
- 4. Consider the 2-way random effects model with interactions:

 $Y_{ijk} = \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \epsilon_{ijk}, i = 1, 2, j = 1, 2, k = 1, \dots, n_{ij},$ 

with variance components  $\sigma_a^2, \sigma_b^2, \sigma_{ab}^2$ , and  $\sigma^2$ , respectively.

- a. Assume  $n_{11} = n_{12} = n_{21} = 2$ , and  $n_{22} = 3$ . Find the Expected Mean Squares as functions of the variance components for the two factors.
- b. Repeat part (a) with  $n_{11} = n_{12} = n_{21} = n_{22} = 2$ , Comment briefly on the results in parts (a) and (b).