# Statistics Qualifying Exam 

12:00 pm - 4:00 pm, Tuesday, August $20^{\text {th }}, 2019$

1. Suppose that random variables $X_{1}$ and $X_{2}$ have the following joint probability mass function (pmf).

|  | $x_{2}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: |
| $x_{1}$ | 0 | 0 | 1 | 2 |

(a) Calculate the expectation of $X_{1}$, i.e., $E\left(X_{1}\right)$.
(b) Let $Y=g\left(X_{1}, X_{2}\right)=X_{2}-X_{1}$. Calculate the expectation of $Y$, i.e., $E\left(g\left(X_{1}, X_{2}\right)\right)$.
(c) Calculate the probability $P\left(X_{1} \geq X_{2}\right)$.
2. Suppose that $X_{1}$ and $X_{2}$ are independent and identically distributed with the exponential distribution whose marginal probability density function (pdf) is given by

$$
f_{X_{j}}\left(x_{j}\right)= \begin{cases}\frac{1}{2} \exp \left(-\frac{x_{j}}{2}\right) & x_{j}>0 \\ 0 & \text { elsewhere }\end{cases}
$$

for $j=1,2$. Define $Y_{1}=\frac{1}{2}\left(X_{1}-X_{2}\right)$ and $Y_{2}=X_{2}$.
(a) Find the joint pdf of $Y_{1}$ and $Y_{2}$.
(b) Show that $Y_{1}$ follows the Laplace distribution (or the double exponential distribution) with zero mean.
3. Let $X_{1}, X_{2}, \ldots, X_{n}$ be a random sample from a population with the pdf as

$$
f_{X}(x)= \begin{cases}1 / \theta, & 0<x<\theta \\ 0, & \text { elsewhere }\end{cases}
$$

Let $Y_{1}<Y_{2}<\ldots<Y_{n}$ be the order statistics. Let $W=Y_{1} / Y_{n}$ and $Q=Y_{n}$.
(a) Find the joint pdf of $W$ and $Q$. Based on the obtained joint distribution funcition of $W$ and $Q$, disucss if $W$ and $Q$ are independent random variables.
(b) Basu's Theorem says that if $T(\mathbf{X})$ is a complete and (minimal) sufficient statistic for $\theta$, then $T(\mathbf{X})$ is independent of every ancillary statistic of $\theta$. Based on Basu's Theorem, discuss if $W$ and $Q$ are independent random variables.
4. Let $X_{1}, \ldots, X_{n}$ be identically and idenpendently distriubted with the $N(0, \theta)$ pdf given as

$$
f(x ; \theta)=\frac{1}{\sqrt{2 \pi \theta}} \exp \left(-\frac{x^{2}}{2 \theta}\right), \quad \theta>0,-\infty<x<+\infty
$$

(a) Derive the maximum likelihood estimator (mle) $\hat{\theta}$ for $\theta$. Show its expectation.
(b) Show that the mle $\hat{\theta}$ obtained in (a) is also the minimum variance unbiased estimator (MVUE) for $\theta$.
(c) We wish to test $H_{0}: \theta=\theta_{0}$ versus $H_{1}: \theta \neq \theta_{0}$. Answer the following questions:
i. Show that the likelihood ratio test depends only on $S=\sum_{i=1}^{n} X_{i}^{2} / \theta_{0}$. That is, $S$ is the test statistic used here.
ii. What is the sampling distribution of $S$ under $H_{0}$ ?
iii. Explicitly state the decision rule of a size $\alpha$ test in the form "reject $H_{0}$ if $S \leq c_{1}$ or $S \geq c_{2}$ " or "reject $H_{0}$ if $S \geq c_{3}$ " with the constants $c_{1}$ and $c_{2}$ or $c_{3}$ clearly specified.
5. Researchers were interested in studying the effect of temperature and light level on the growth of bacterial colonies on potato leaflets. Bacteria were inoculated onto a total of 48 leaflets. The leaflets were randomly assigned to treatment with one of four temperatures (10, 15, 20, or $25^{\circ} \mathrm{C}$ ) and one of three light levels ( $\mathrm{A}=$ low, $\mathrm{B}=$ medium, or $\mathrm{C}=$ high). Four weeks after inoculation, the $\log$ of the area of the bacterial colony on each leaflet was measured as the response variable. A completely randomized design was used with four leaflets for each combination of temperature and light intensity. Use the SAS code and output on the next page of your exam to answer the following questions.
(a) Were there significant differences among the 12 treatment means? Give an appropriate test statistic, its degrees of freedom, the $p$-value, and a brief conclusion.
(b) Were there any significant differences among the temperature lsmeans? Give an appropriate test statistic, its degrees of freedom, the $p$-value, and a brief conclusion.
(c) Two models have been fit to the data. Does the second model fit the data adequately? Give an appropriate test statistic, its degrees of freedom, the $p$-value, and a brief conclusion.

Now, for each of the three light intensities, suppose there is a linear relationship between the mean of the response variable and temperature.
(d) For low light level A, provide the estimated linear regression equation relating mean response to temperature.
(e) For high light level C, give an $95 \%$ confidence interval for the slope of the linear regression equation. Based on this confidence interval, is there evidence that temperature affected bacterial colony growth at high light level C? Explain.

| $d f$ | 41 | 42 | 43 | 44 | 45 |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $t_{0.05 ; d f}$ | 1.6829 | 1.6820 | 1.6811 | 1.6802 | 1.6794 |
| $t_{0.025 ; d f}$ | 2.0195 | 2.0181 | 2.0167 | 2.0154 | 2.0141 |

(f) Estimate the difference between the slope for low light level A and the slope for high light level C, and determine if that difference is significantly different from 0 . Provide the estimated difference, a test statistic, its degrees of freedom, the $p$-value, and a brief conclusion.

```
proc glm;
    class light temp;
    model y=light temp light*temp;
run;
```

|  | $C l a s s$ | Level Information |  |
| :--- | :---: | :---: | :---: | :--- |
| Class | Levels | Values |  |
| light | 3 | A B C |  |
| temp | 4 | $10 \quad 15 \quad 20$ | 25 |
| Number of observations | 48 |  |  |


| Source | DF | Sum of Squares | Mean Square | F | Value | Pr $>$ F |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mbdel | 11 | 2420.799306 | 220.072664 |  | 8.50 | <. 0001 |
| Error | 36 | 932.134425 | 25.892623 |  |  |  |
| Corrected Total | 47 | 3352.933731 |  |  |  |  |
| Source | DF | Type I SS | Mean Square | F | Value | Pr $>\mathrm{F}$ |
| 1 ight | 2 | 1588.672888 | 794.336444 |  | 30.68 | <. 0001 |
| temp | 3 | 441.252123 | 147.084041 |  | 5.68 | 0.0027 |
| $1 \mathrm{ight} * \mathrm{temp}$ | 6 | 390.874296 | 65.145716 |  | 2.52 | 0.0389 |
| Source | DF | Type III SS | Mean Square | F | Value | Pr $>\mathrm{F}$ |
| 1 ight | 2 | 1588.672887 | 794.336444 |  | 30.68 | <. 0001 |
| temp | 3 | 441.252123 | 147.084041 |  | 5.68 | 0.0027 |
| 1ight * emp | 6 | 390.874296 | 65.145716 |  | 2. 52 | 0.0389 |

```
proc glm;
        class light;
    model y=light temp light*temp / solution;
run;
```

| Cl ass |  | LevelInformation <br> Class |
| :--- | :---: | :---: |
| Levels | Values |  |
| 1 ight | 3 | A B C |
| Number of observations | 48 |  |


| Sum of |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Source D |  |  | DF | Squares | Mean Square | F | Value |  |
| Mbdel |  |  | 5 | 2224. 265059 | 444.853012 |  | 16. 55 |  |
| Error |  |  | 42 | 1128.668673 | 26.873064 |  |  | $\text { <. } 0001$ |
| Corrected Total 4 |  |  | 47 | 3352.933731 |  |  |  |  |
| Source |  |  | DF | Type I SS | Mean Square |  | Value | Pr $>\mathrm{F}$ |
| 1 ight |  |  | 2 | 1588.672888 | 794.336444 |  | 29.56 | <. 0001 |
| t emp |  |  | 1 | 373.276984 | 373.276984 |  | 13.89 | 0.0006 |
| temp*light |  |  | 2 | 262.315188 | 131.157594 | 4.88 |  | 0.0124 |
| Source |  |  | DF | Type III SS | Mean Square |  | Val ue | Pr $>\mathrm{F}$ |
| 1 ight |  |  | 2 | 12.3292407 | 6. 1646204 |  | 0.23 | 0.7960 |
| t emp |  |  | 1 | 373.2769837 | 373.2769837 |  | 13.89 | 0.0006 |
| temp*light |  |  | 2 | 262.3151875 | 131.1575938 |  | 4.88 | 0.0124 |
|  | St andard |  |  |  |  |  |  |  |
| Parameter |  | Estimate |  | Error | t Value | $\operatorname{Pr}>\|t\|$ |  |  |
| Intercept |  | 2.044500000 |  | 4.25902782 | 0.48 | 0.6337 |  |  |
| 1 ight | A | -2.307500000 |  | 6.02317490 | -0.38 | 0.7036 |  |  |
| 1 ight | B | 1.760000000 |  | 6.02317490 | 0.29 | 0.7716 |  |  |
| lighttemp |  | 0.000000000 |  | . | . |  |  |  |
|  |  | 0. 276600000 |  | 0.23183211 | 1. 19 | 0.2395 |  |  |
| temp*1ight | A | 0.808000000 |  | 0.32786011 | 2.46 | 0.0179 |  |  |
| temp*light | B | -0.141250000 |  | 0. 32786011 | -0.43 | 0.6 | 688 |  |
| temp*1ight | C | 0.000000000 |  | . | . |  |  |  |

6. The following is part of ANOVA table for a simple linear regression model

$$
Y_{i}=\beta_{0}+\beta_{1} X_{i}+\varepsilon_{i}
$$

where $\varepsilon_{i}$ 's are i.i.d. from $N\left(0, \sigma^{2}\right), i=1, \ldots, n$, and $n$ is the number of observations. (No need to complete the table.)

| Sum of |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Source | DF | Squares | Mean Square | F Value | $\mathrm{Pr}>\mathrm{F}$ |
| Model | *** | * | 252378 | 105.88 | <. 0001 |
| Error | 23 | ****** | ****** |  |  |
| Corrected Total | *** | ****** |  |  |  |

(a) Compute the coefficient of determination $R^{2}$.
(b) Assume that we now know the least square estimate of $\beta_{1}$ is $b_{1}=3.57$. Construct a twosided $t$-test of whether or not $\beta_{1}=3$. State the null and alternative hypotheses, the value of the test statistic, the sampling distribution under the null hypothesis and the decision rule. [Note: You don't need to calculate the p-value or make the conclusion.]
7. An experiment was conducted to determine the effects of four different pesticides on the yield of fruit from three different varieties of a citrus tree. Eight trees from each variety were available and the the four pesticides were then randomly assigned to two trees of each variety. Yields of fruit (in bushels per tree) were obtained after the test period. The mean yields for each combination of pesticide and variety are given below.

|  | Variety |  |  |  |
| :---: | :--- | :--- | :--- | ---: |
| Pesticide | 1 | 2 | 3 | Pesticide Means |
| 1 | 44 | 48 | 67 | 53.00 |
| 2 | 52.5 | 62.5 | 88.5 | 67.83 |
| 3 | 40.5 | 47.5 | 65.5 | 51.17 |
| 4 | 50.5 | 79 | 92 | 73.83 |
| Variety Means | 46.875 | 59.25 | 78.25 |  |

Suppose the following statistical model is used to fit the data.

$$
Y_{i j k}=\mu+\tau_{i}+\beta_{j}+(\tau \beta)_{i j}+\epsilon_{i j k} ; k=1,2
$$

where $\tau_{i}(i=1,2,3,4)$ and $\beta_{j}(j=1,2,3)$ are the effects of pesticide and variety, and $(\tau \beta)_{i j}$ are their interactions. For parameter estimation, we impose the following constraints as in the lecture notes: $\sum_{i} \tau_{i}=\sum_{j} \beta_{j}=\sum_{i}(\tau \beta)_{i j}=\sum_{j}(\tau \beta)_{i j}=0$.
Some SAS output is shown.

```
The GLM Procedure
Dependent Variable: yield
Source Sum of Squares
Model 6680.458333
Error 507.500000
Corrected Total 7187.958333
Source Type I SS
pesticide 2227.458333
variety 3996.083333
pesticide*variety 456.916667
```

(a) What are the estimates of $\tau_{3}$ and $(\tau \beta)_{23}$ ?
(b) Provide the degrees of freedom corresponding to each of the sums of squares in the output, which are marked by "???" below.

| Source | Degrees of freedom |
| :--- | ---: |
| Model | $? ? ?$ |
| Error | $? ? ?$ |
| Corrected Total | $? ? ?$ |
|  |  |
|  |  |
| pesticide | $? ? ?$ |
| variety | $? ? ?$ |
| pesticide * variety | $? ? ?$ |

(c) Do the effects of the pesticides on yield dependent on the variety of citrus tree? Conduct an appropriate test to answer this question. To get full credits, give hypotheses, a test statistic, determine its degrees of freedom, use an appropriate value from the table below, state the p-value (or its range) and give a conclusion using $\alpha=0.05$.

| $\left(d f_{1}, d f_{2}\right)$ | $(5,10)$ | $(5,11)$ | $(5,12)$ | $(5,13)$ | $(5,14)$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $F_{0.055} d f_{1}, d f_{2}$ | 3.3258 | 3.2039 | 3.1059 | 3.0254 | 2.9582 |
| $\left(d f_{1}, d f_{2}\right)$ | $(6,10)$ | $(6,11)$ | $(6,12)$ | $(6,13)$ | $(6,14))$ |
| $F_{0.05 ; d f_{1}, d f_{2}}$ | 3.2172 | 3.0946 | 2.9961 | 2.9153 | 2.8477 |

(d) Use Tukey's method to perform a pairwise comparison for different varieties. Report the critical difference and report your results of comparison (using $\alpha=0.05$ ). You can report the result as we have done in class by labeling significantly different combinations with different Latin letters.

| $\left(d f_{1}, d f_{2}\right)$ | $(2,10)$ | $(2,11)$ | $(2,12)$ | $(2,13)$ | $(2,14)$ |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $q_{0.025 ; d f_{1}, d f_{2}}$ | 3.7247 | 3.6672 | 3.6204 | 3.5817 | 3.5491 |
| $q_{0.05 ; d f_{1}, d f_{2}}$ | 3.1511 | 3.1127 | 3.0813 | 3.0552 | 3.0332 |
| $\left(d f_{1}, d f_{2}\right)$ | $(3,10)$ | $(3,11)$ | $(3,12)$ | $(3,13)$ | $(3,14)$ |
| $q_{0.025 ; d f_{1}, d f_{2}}$ | 4.4740 | 4.3913 | 4.3243 | 4.2687 | 4.2220 |
| $q_{0.05 ; d f_{1}, d f_{2}}$ | 3.8768 | 3.8196 | 3.7729 | 3.7341 | 3.7014 |

(e) Suppose Pesticides 1 and 2 are sold by Company A and Pesticides 3 and 4 are sold by Company B. Conduct a test or construct a confidence interval that can be used to compare the effectively of Company A's pesticides to the effectiveness of Company B's pesticides. Show your work and provide a conclusion.

| $\left(d f_{1}, d f_{2}\right)$ | $(1,10)$ | $(1,11)$ | $(1,12)$ | $(1,13)$ | $(1,14)$ |
| :--- | :---: | :---: | :---: | :---: | :---: |
| $F_{0.05 ; d f_{1}, d f_{2}}$ | 4.9646 | 4.8443 | 4.7472 | 4.6672 | 4.6001 |
| $d f$ | 10 | 11 | 12 | 13 | 14 |
| $t_{0.05 ; d f}$ | 1.8125 | 1.7959 | 1.7823 | 1.7709 | 1.7613 |
| $t_{0.025 ; d f}$ | 2.2281 | 2.2010 | 2.1788 | 2.1604 | 2.1448 |

8. You have been asked to design an experiment to compare the effect of two types of plant food on tomato plant growth using sixteen plants. These sixteen plants will be grown on a greenhouse bench in two rows of eight as shown below ( 2 rows of 8 plants). Even though this is in a greenhouse, you are concerned about air temperature (a nuisance factor) affecting plant growth. For each item below,

- describe how you would allocate the treatments to these sixteen tomato plants;
- give the model, and write out the ANOVA table in terms of sources and degrees of freedom.
(a) You are told that there is no temperature gradient.

|  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |  |  |  |

(b) You are told that there is a considerable temperature gradient running against the rows $(\downarrow)$.

(c) You are told there is a moderate temperature gradient that runs along the bench $(\leftarrow)$.

|  |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  |  |  |  |  |  |  |  |

