

# Statistics and Probability Prelim Exam

9am-1pm, Wednesday, May 4, 2016

Note: To pass this exam, you need to pass both Statistics and Probability parts

## Statistics Part

- Suppose  $X_1, \dots, X_n$  are iid with pdf  $f(x|\theta) = \theta e^{-\theta x}$ ,  $x > 0$ , where  $\theta > 0$  is unknown. Assume a prior distribution given by  $\pi(\theta) = e^{-\theta}$ ,  $\theta > 0$ .
  - Find the posterior distribution of  $\theta$  and the posterior mean.
  - Now, assume a loss function for estimating  $\theta$  given by  $L(\theta, a) = e^{(a-\theta)} - (a - \theta) - 1$ . Find the Bayes rule for estimating  $\theta$  based on this loss function.
- Let  $X_1, \dots, X_n$  be a random sample from  $Poisson(\mu)$  distribution. Suppose we want to estimate  $\theta = P(X_1 = 0) = e^{-\mu}$ . Let  $T_1$  be the maximum likelihood estimator of  $\theta$ , and let  $T_2 = (1/n) \sum_i I_{\{0\}}(X_i)$  where  $I_{\{0\}}(x)$  is the indicator function which is equal to 1 if  $x = 0$ , and is equal to 0 otherwise. Note that  $T_2$  is the proportion of zeroes in the sample.
  - Find  $T_1$  and show that  $T_1$  and  $T_2$  are each asymptotically normally distributed.
  - Find the asymptotic relative efficiency of  $T_2$  with respect to  $T_1$ .
  - Determine if either or both of  $T_1$  and  $T_2$  are unbiased for  $\theta$ .
- Suppose that  $X_1, \dots, X_n$  are iid  $Poisson(\lambda)$  where  $0 < \lambda < \infty$  is unknown.
  - Show that  $T(X_1, \dots, X_n) = \sum_{i=1}^n X_i$  is a complete sufficient statistic for  $\lambda$ .
  - Find the Uniformly Minimum Variance Unbiased Estimator (UMVUE) of  $\tau_1(\lambda) = \lambda^2$ .
  - Find the UMVUE of  $\tau_2(\lambda) = \lambda e^{-\lambda}$ , that is  $P(X_1 = 1) = \lambda e^{-\lambda}$ .
- Let  $X_1, \dots, X_n$  be a sample from the *inverse Gaussian* distribution  $I(\mu, \tau)$  with density

$$\sqrt{\frac{\tau}{2\pi x^3}} \exp\left(-\frac{\tau}{2x\mu^2}(x - \mu)^2\right), \quad x > 0, \quad \tau, \mu > 0.$$

- Find the moment generating function of  $X_1$ .
- Show that  $V = \frac{\tau}{\bar{X}\mu^2}(X - \mu)^2 \sim \chi_1^2$
- Show that  $\bar{X} = \sum_{i=1}^n X_i/n \sim I(\mu, n\tau)$ .
- Show that there exists a UMP test for testing  $H_0 : \mu \leq \mu_0$  versus  $H_1 : \mu > \mu_0$  when  $\tau$  is known.
- Show that there exists a UMP test for testing  $H_0 : \tau \leq \tau_0$  versus  $H_1 : \tau > \tau_0$  when  $\mu$  is known.

## Probability Part

5. Prove the following identity: for all  $p \geq 1$  and any random variable  $X$ ,

$$\mathbb{E}|X|^p = p \int_0^\infty x^{p-1} \mathbb{P}(|X| > x) dx.$$

6. Consider i.i.d. Poisson random variables  $\{X_n\}_{n \in \mathbb{N}}$  with parameter  $\lambda > 0$ . Consider  $Y_n := X_n X_{2n}, n \in \mathbb{N}$  and  $T_n := Y_1 + \dots + Y_n$ .

- (a) Compute  $\mathbb{E}T_n$ .
- (b) Find an explicit constant  $C$  such that  $\text{Var}(T_n) \leq Cn$  for all  $n \in \mathbb{N}$ . Explain clearly how you determine the constant  $C$ . The constant does not have to be optimal and may depend on  $\lambda$ . However, it must be independent from  $n$ .
- (c) Find a sequence of real numbers  $\{a_n\}_{n \in \mathbb{N}}$ , such that

$$\frac{T_n}{a_n} \rightarrow 1 \text{ in probability.}$$

Justify your result.

7. Let  $X_1, X_2, \dots$  be independent non-negative random variables. Suppose that there exists constants  $K < \infty$  such that  $\sup_{n \in \mathbb{N}} \mathbb{E} \exp(X_n) \leq K$ .

- (a) Show that for all  $x > 0, n \in \mathbb{N}$ ,

$$\mathbb{P}(X_n > x) \leq K e^{-x}.$$

- (b) Show that for all  $\epsilon > 0$ ,

$$\mathbb{P}\left(\frac{X_n}{\log n} > 1 + \epsilon \text{ i.o.}\right) = 0.$$

You may use the inequality in part (a) directly.

8. (a) State Lindeberg–Feller central limit theorem.
- (b) Use Lindeberg–Feller central limit theorem to prove the following. Given a sequence  $\{X_n\}_{n \in \mathbb{N}}$  of independent but *not necessarily identically distributed* random variables with mean zero, variance 2 and  $\mathbb{E}|X_k|^4 < 17$ , show that the central limit theorem for the partial sums  $S_n := X_1 + \dots + X_n$  holds in the form of

$$\frac{S_n - b_n}{a_n} \Rightarrow \mathcal{N}(0, 1),$$

and specify the normalizing constants  $\{a_n\}_{n \in \mathbb{N}}$  and  $\{b_n\}_{n \in \mathbb{N}}$ .